

ÉCOLE POLYTECHNIQUE FÉDÉRALE DE LAUSANNE

School of Computer and Communication Sciences

Foundations of Data Science
Fall 2025

Assignment date: Wednesday, November 12th, 2025, 13:15
Due date: Wednesday, November 12th, 2025, 14:45

Midterm Exam – AAC 231

This exam is open book. No electronic devices of any kind are allowed. There are three problems. Good luck!

Only answers given on this handout count.

Name: _____

SCIPER: _____

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Recall: The Beta distribution with parameters α and β is the probability density function given by

$$p_{\alpha,\beta}(x) = \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha,\beta)}$$

where $B(\alpha,\beta)$ is the beta function defined as

$$B(\alpha,\beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}.$$

The mean of a random variable X with probability density function $p_{\alpha,\beta}(x)$ is $\mathbb{E}[X] = \frac{\alpha}{\alpha+\beta}$.

Problem 1 (DPI for TV distance – 8 pts). Show that the Total Variation Distance satisfies the Data Processing Inequality. That is, using the definitions given in Lemma 4.3 of the Lecture Notes: a probability kernel $W(y|x)$, an original distribution P , and the distribution $\tilde{P}$ whose probability mass function is $\tilde{p}(y) = \sum_x W(y|x)p(x)$, show that

$$\delta(\tilde{P}, \tilde{Q}) \leq \delta(P, Q),$$

where $\delta(P, Q)$ denotes the Total Variation distance between P and Q .

Remark: If you refer to class materials, be precise (Theorem or equation numbers, Homework problem identifiers and so on.) Your overall argument must be complete.

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Problem 2 (Thompson Sampling for Two-Armed Bernoulli Bandits – 12 pts). Consider the following two-arm setting:

- Let μ_1 and μ_2 be two independent samples from the uniform distribution on the interval $[0, 1]$. As in class, denote $\Delta = |\mu_1 - \mu_2|$. But note that we do not know which arm is better.
- Now, pulling arm $a \in \{1, 2\}$ at round t yields a reward

$$X_t \sim \text{Bernoulli}(\mu_a).$$

- Let $S_a(t-1)$ and $F_a(t-1)$, respectively, denote the number of observed successes (reward=1) and failures (reward=0) of arm a up to round $t-1$ (for $a=1$ and $a=2$).

We use the following algorithm (*Thompson Sampling*):

At round $t = 1, 2, \dots$:

(a) Sample independently

$$\tilde{\mu}_1^{(t)} \sim \text{Beta}(1 + S_1(t-1), 1 + F_1(t-1))$$

$$\tilde{\mu}_2^{(t)} \sim \text{Beta}(1 + S_2(t-1), 1 + F_2(t-1))$$

Note: See the exam's cover page for a reminder about the Beta distribution.

(b) Pull the arm with the larger sample:

$$A_t = \arg \max_{a \in \{1, 2\}} \tilde{\mu}_a^{(t)}.$$

(c) Observe reward $X_t \in \{0, 1\}$ from arm A_t .

As in class, let $T_{\text{bad}}(n)$ denote the number of times the bad arm (the arm with the smaller value μ_a) is pulled up to horizon n . As in class, the regret can then be written as $R_n = \Delta \mathbb{E}[T_{\text{bad}}(n)]$.

Tasks. Answer the following:

- (a) [2 points] (**Formulation**) Recall that the means μ_a were drawn from a uniform prior distribution. Suppose by time t , arm a has been pulled $T_a(t)$ times with $S_a(t)$ successes and $F_a(t)$ failures. Give an expression for the posterior distribution $p(\mu_a | S_a(t), F_a(t))$. What type of a distribution is it?

- (b) [2 points] (**Partition of the time axis**) Show that for $m > 0$

$$\mathbb{E}[T_a(n)] \leq m + \mathbb{E}[\# \text{ of pulls of arm } a \text{ after it has already been pulled } m \text{ times}].$$

(Hint: decompose according to whether $T_a(t) < m$ or $T_a(t) \geq m$, where $T_a(t)$ is the number of pulls of arm a before round t .)

- (c) [3 points] (**Concentration of the Beta posterior**) *This part is a little more tedious. You may want to save it for last and first tackle parts (d) and (e).* For a constant $c > 0$ to be chosen later, define

$$m := \left\lceil \frac{c \log n}{\Delta^2} \right\rceil.$$

Argue that once arm a has been pulled m times, the posterior distribution you found in Part (a) is concentrated around the true μ_a , in the sense that for any $\varepsilon > 0$, there exist constants $c_1, c_2 > 0$ such that, for all t with $T_a(t) \geq m$,

$$\begin{aligned} \mathbb{P}(\tilde{\mu}_a^{(t)} > \mu_a + \varepsilon \mid H_{t-1}) &\leq c_1 \exp(-c_2 T_a(t) \varepsilon^2), \\ \mathbb{P}(\tilde{\mu}_a^{(t)} < \mu_a - \varepsilon \mid H_{t-1}) &\leq c_1 \exp(-c_2 T_a(t) \varepsilon^2), \end{aligned}$$

where as in class, we are using H_{t-1} to denote the entire history up to time $t-1$, that is, $H_{t-1} = (A_1, X_1, A_2, X_2, \dots, A_{t-1}, X_{t-1})$.

Hint: Calculate $\hat{\mu}_{a,t} := \mathbb{E}[\tilde{\mu}_a^{(t)} \mid H_{t-1}]$. Then, as in class, split by conditioning onto the events G_a and G_a^c , where $G_a := \{|\hat{\mu}_{a,t} - \mu_a| \leq \frac{\varepsilon}{4}\}$.

- (d) [3 points] (**Bounding the selection of the suboptimal arm**) Use the previous item with $\varepsilon = \Delta/2$. For large enough t (when both arms have been pulled m times), provide an upper bound for the probability that we select the bad arm,

$$\mathbb{P}(A_t = \text{bad} \mid H_{t-1})$$

in terms of some constant $c' > 0, \Delta$ and n .

- (e) [2 points] (**Final bound on $\mathbb{E}[T_{\text{bad}}(n)]$**) Pick a large enough c and conclude that

$$\mathbb{E}[T_{\text{bad}}(n)] \leq \frac{c \log n}{\Delta^2} + O(1),$$

and hence the regret of Thompson Sampling in this two-armed Bernoulli bandit is bounded by

$$R_n = \Delta \mathbb{E}[T_{\text{bad}}(n)] = O\left(\frac{\log n}{\Delta}\right).$$

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Problem 3 (Fisher goes Beta – 10 pts). In this problem, we will establish a few properties of the $\text{Beta}(\alpha, m)$ distribution when m is an integer. See the exam's cover page for a reminder about the Beta distribution.

Hint: Recall from Homework 4 that for any $x > 0$, we have $x\Gamma(x) = \Gamma(x+1)$, and that $\Gamma(0) = 1$.

- (a) [3 points] Let $X \sim \text{Beta}(\alpha, m)$, where m is a positive integer. For $\alpha > 0$, find

$$\mathbb{E}[\log X].$$

Hint: Try finding $\mathbb{E}_{\alpha, m} \left[\frac{d}{d\alpha} \log p_{\alpha, m}(X) \right]$. If finding the expectation for a general m seems too difficult, try $m = 1, 2$ first.

- (b) [3 points] Show that, for any parametric family of density functions $p_\theta(x)$ such that the support is independent of θ , the Fisher information

$$I(\theta) = -\mathbb{E}_\theta \left[\frac{d^2}{d\theta^2} \log p_\theta(X) \right].$$

- (c) [4 points] Using part (b), show that for $X \sim \text{Beta}(\alpha, m)$ for a fixed integer $m > 0$, whenever $\alpha \in (0, 1)$,

$$I(\alpha) \leq \frac{1}{\alpha^2} + \frac{\pi^2}{6}.$$

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