

## Solutions to Homework 9

**Exercise 1.**

a) Fix  $\epsilon > 0$ . Then

$$\mathbb{P}(|X_n - C| > \epsilon) = \mathbb{P}(X_n < C - \epsilon) + \mathbb{P}(X_n > C + \epsilon) \leq F_{X_n}(C - \epsilon) + (1 - F_{X_n}(C + \epsilon)) \xrightarrow{n \rightarrow \infty} 0$$

whenever  $X_n \xrightarrow[n \rightarrow \infty]{d} C$ .

b) Assume that  $r > s \geq 1$  and that

$$\mathbb{E}(|X_n - X|^r) \xrightarrow{n \rightarrow \infty} 0.$$

Similar to exercise 4 a) on the midterm, we can use Jensen's inequality to show

$$\mathbb{E}(|X_n - X|^s)^{\frac{r}{s}} \leq \mathbb{E}(|X_n - X|^r)$$

by applying the convex function  $f(x) = x^{r/s}$ . Then

$$\mathbb{E}(|X_n - X|^s)^{\frac{r}{s}} \leq \mathbb{E}(|X_n - X|^r) \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \mathbb{E}(|X_n - X|^s)^{\frac{r}{s}} \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \mathbb{E}(|X_n - X|^s) \xrightarrow{n \rightarrow \infty} 0.$$

And therefore  $X_n \xrightarrow[n \rightarrow \infty]{L^s} X$ .

c) We have that

$$|\mathbb{E}(X_n) - \mathbb{E}(X)| = |\mathbb{E}(X_n - X)| \leq \mathbb{E}(|X_n - X|) \xrightarrow{n \rightarrow \infty} 0.$$

Therefore  $\mathbb{E}(X_n) \xrightarrow[n \rightarrow \infty]{} \mathbb{E}(X)$ .

The converse is not true. Consider the sequence  $(X_n, n \geq 1)$  of i.i.d. Bernoulli( $p$ ) random variables, with  $0 < p < 1$ . Then  $\mathbb{E}(X_n) \xrightarrow[n \rightarrow \infty]{} \mathbb{E}(X)$ . However

$$\mathbb{E}(|X_n - X|) = \frac{1}{2} \cdot 1 + \frac{1}{2} \cdot 0 = \frac{1}{2}.$$

This does not converge to zero as  $n$  goes to infinity.

**Exercise 2.** a) Let us compute first

$$\mathbb{E}(S_1) = \frac{1}{2} \left( \frac{3S_0}{2} + \frac{S_0}{2} \right) = S_0$$

Assuming now that  $\mathbb{E}(S_n) = S_0$  (more precisely, that the expectation stays constant over  $n$  coin tosses), let us compute  $\mathbb{E}(S_{n+1})$ :

$$\begin{aligned} \mathbb{E}(S_{n+1}) &= \mathbb{E}(S_{n+1} | \{X_1 = +1\}) \mathbb{P}(\{X_1 = +1\}) + \mathbb{E}(S_{n+1} | \{X_1 = -1\}) \mathbb{P}(\{X_1 = -1\}) \\ &= \frac{1}{2} \left( \mathbb{E}(S_{n+1} | \{S_1 = \frac{3S_0}{2}\}) + \mathbb{E}(S_{n+1} | \{S_1 = \frac{S_0}{2}\}) \right) = \frac{1}{2} \left( \frac{3S_0}{2} + \frac{S_0}{2} \right) = S_0 \end{aligned}$$

*Note:* The computation is slightly unorthodox here, but we will see a cleaner way to prove this later in the course.

b)  $Y_n$  is the sum of  $n$  i.i.d. random variables, as the following computation shows:

$$Y_n = \log\left(\frac{S_n}{S_0}\right) = \log\left(\prod_{j=1}^n \left(1 + \frac{X_j}{2}\right)\right) = \sum_{j=1}^n \log\left(1 + \frac{X_j}{2}\right)$$

and these random variables are bounded, so by the central limit theorem,

$$\frac{Y_n - n\mu}{\sqrt{n}\sigma} \xrightarrow[n \rightarrow \infty]{d} Z \sim \mathcal{N}(0, 1)$$

where  $\mu = \mathbb{E}(\log(1 + X_1/2)) = \frac{1}{2}(\log(3/2) + \log(1/2)) \simeq -0.144$  and

$$\sigma^2 = \text{Var}(\log(1 + X_1/2)) = \frac{1}{2}(\log(3/2)^2 + \log(1/2)^2) - \mu^2 \simeq 0.3$$

This is saying that for large  $n$ , we have

$$Y_n \simeq -0.144n + \sqrt{0.26n} Z \quad \text{in particular: } Y_{100} \simeq -14.4 + 5.4 Z$$

Therefore

$$\begin{aligned} \mathbb{P}(\{S_{100} > S_0/10\}) &= \mathbb{P}(\{S_{100}/S_0 > 1/10\}) = \mathbb{P}(\{Y_{100} > -\log(10)\}) \\ &\simeq \mathbb{P}\left(\left\{Z > \frac{-2.3 + 14.4}{5.4}\right\}\right) = \mathbb{P}(\{Z > 2.24\}) \end{aligned}$$

which is roughly 1% (so you can imagine what  $\mathbb{P}(\{S_{100} > S_0\})$  looks like ...).

Therefore, the process  $(S_n, n \geq 1)$ , unexpectedly perhaps, “crashes” to zero with high probability as  $n$  gets large, even though it seemed a priori a “fair game” with constant expectation. This is an important example among a large class of processes called “martingales”; we will come back to this!

*Note:* The random process  $(S_n, n \geq 1)$  is not unrelated to the following *deterministic* process defined recursively as

$$x_0 \in \mathbb{N}^*, \quad x_{n+1} = \begin{cases} x_n/2 & \text{if } x_n \text{ is even} \\ 3x_n + 1 & \text{if } x_n \text{ is odd} \end{cases}$$

in which an even number gets multiplied by  $1/2$  and an odd number gets approximately multiplied by  $3/2$  (because it first gets multiplied by  $3$  and then necessarily divided by  $2$ , as  $3x_n + 1$  is even). So if you consider that even and odd numbers appear naturally with probability  $1/2$ , then the two processes have something in common. But in the deterministic case, one has no proof that the process ultimately reaches the value  $1$  as  $n$  gets large: this is the famous Collatz conjecture, which remains unsolved until now.

**Exercise 3.** a) let us compute  $\mathbb{E}(S_n) = \sum_{j=1}^n \mathbb{E}(X_j^{(n)}) = n \frac{\lambda}{n} = \lambda$  and

$$\text{Var}(S_n) = \sum_{j=1}^n \text{Var}(X_j^{(n)}) = n \frac{\lambda}{n} \left(1 - \frac{\lambda}{n}\right) = \lambda - \frac{\lambda^2}{n}$$

b) So  $\mu = \lim_{n \rightarrow \infty} \mathbb{E}(S_n) = \lambda$  and  $\sigma^2 = \lim_{n \rightarrow \infty} \text{Var}(S_n) = \lambda$ .

c) Let us compute the characteristic function of  $S_n$ :

$$\begin{aligned}\phi_{S_n}(t) &= \mathbb{E}(\exp(itS_n)) = \mathbb{E}(\exp(it(X_1^{(n)} + \dots + X_n^{(n)}))) = \mathbb{E}(\exp(itX_1^{(n)})) \cdots \mathbb{E}(\exp(itX_n^{(n)})) \\ &= \left(\mathbb{E}(\exp(itX_1^{(n)}))\right)^n = \left(e^{it\frac{\lambda}{n}} + 1 - \frac{\lambda}{n}\right)^n = \left(1 + \frac{\lambda(e^{it} - 1)}{n}\right)^n \xrightarrow{n \rightarrow \infty} \exp(\lambda(e^{it} - 1))\end{aligned}$$

This limiting function is the characteristic function of  $Z \sim \mathcal{P}(\lambda)$ . Indeed, one can check that

$$\phi_Z(t) = \mathbb{E}(\exp(itZ)) = \sum_{k \geq 0} e^{itk} \frac{\lambda^k e^{-\lambda}}{k!} = e^{-\lambda} \sum_{k \geq 0} \frac{(\lambda e^{it})^k}{k!} = \exp(\lambda(e^{it} - 1))$$

which allows us to conclude that  $S_n \xrightarrow[n \rightarrow \infty]{d} Z$ .

d) The computation of the characteristic function is similar here:

$$\mathbb{E}(e^{itT_n}) = \left(\frac{1}{n} e^{it} + \left(1 - \frac{1}{n}\right)\right)^{\lceil \lambda n \rceil} = \left(1 + \frac{1}{n}(e^{it} - 1)\right)^{\lceil \lambda n \rceil} \xrightarrow{n \rightarrow \infty} \exp(\lambda(e^{it} - 1))$$

and leads actually exactly to the same result:  $T_n$  converges in distribution towards a Poisson random variable  $Z$  of parameter  $\lambda$ .

e) No, as each random variable  $S_n$  is constructed from a different set of random variables  $X_1^{(n)}, \dots, X_n^{(n)}$ , which depends on  $n$ . The same holds for the random variables  $T_n$ .